Hierarchical Bayesian Sparsity: ℓ_2 Magic.

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Linear inverse problem

Estimate $x \in \mathbb{R}^n$ with an underdetermined observation model

$$b = Ax + e, \quad A \in \mathbb{R}^{m \times n},$$

where

- *m* ≪ *n*
- additive Gaussian noise e is a realization of random variable $E \sim \mathcal{N}(0, I_m)$
- a priori, x is believed to be sparse, that is

 $\|x\|_0 \ll n.$

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Sparsity promotion: hierarchical model

• Hierarchical conditionally Gaussian prior hypermodel

$$X \sim \mathcal{N}(0, \mathsf{D}_{\theta}), \quad \mathsf{D}_{\theta} = \operatorname{diag}(\theta_1, \ldots, \theta_n),$$

• Assume the prior variances $\theta_j > 0$ are mutually independent random variables following a Gamma distribution,

$$\Theta_j \sim \operatorname{Gamma}(eta, heta_j^*) \propto heta_j^{eta-1} \exp\left(-rac{ heta_j}{ heta_j^*}
ight), \quad 1 \leq j \leq n.$$

Posterior density

$$\pi_{\boldsymbol{X},\boldsymbol{\Theta}|\boldsymbol{B}}(\boldsymbol{x},\boldsymbol{\theta}) \propto \exp\left(-\frac{1}{2}\|\boldsymbol{b} - \boldsymbol{A}\boldsymbol{x}\|^2 - \frac{1}{2}\sum_{j=1}^n \frac{x_j^2}{\theta_j} + \eta \sum_{j=1}^n \log \theta_j - \sum_{j=1}^n \frac{\theta_j}{\theta_j^*}\right)$$

where $\eta = \beta - 3/2 > 0$.

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Iterared Alternating Sequential (IAS) algorithm

To compute x_{MAP} we minimize the Gibbs energy

$$\mathscr{E}(x;\theta) = \underbrace{\frac{1}{2} \|b - Ax\|^2}_{(b)} + \underbrace{\sum_{j=1}^n \frac{x_j^2}{2\theta_j}}_{(b)} - \sum_{j=1}^n \left(\eta \log \theta_j - \frac{\theta_j}{\theta_j^*}\right)_{(b)}$$
(1)

Given the initial value $\theta^0 = \theta^*$, $x^0 = 0$, and k = 0, iterate until convergence: (a) Update $x^k \to x^{k+1}$ by minimizing $\mathscr{E}(x; \theta^k)$; (b) Update $\theta^k \to \theta^{k+1}$ by minimizing $\mathscr{E}(x^{k+1}; \theta)$; (c) Increase $k \to k+1$.

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Exact IAS algorithm

Initialize: k = 0, $\theta_0 = \theta^*$; While $\|\theta_k - \theta_{k-1}\| > \text{tol}$ • Update x; $x_{k+1} = \operatorname{argmin} \left\{ \|b - Ax\|^2 + \|D_{\theta}^{-1/2}x\|^2 \right\}$ by solving $\begin{bmatrix} A \\ D_{\theta}^{-1/2} \end{bmatrix} x = \begin{bmatrix} b \\ 0 \end{bmatrix}$

in the least squares sense.

2 Setting $x = x_{k+1}$, update the components of θ_{k+1} according to the formula

$$\theta_j = \theta_j^* \left(\frac{\eta}{2} + \sqrt{\frac{\eta^2}{4} + \frac{x_j^2}{2\theta_j^*}} \right)$$

Convexity and Convergence of exact IAS

The exact IAS algorithms with the Gamma hyperprior is such that:

- The Gibbs energy functional is strictly convex
- The Gibbs functional has a unique minimizer
- If Step 1 and Step 2 are solved exactly the algorithm converges to the global minimizer
- For $\eta > 0$ small, the Gibbs energy (1) is approximately equal to the penalized least squares functional with a weighted ℓ_1 -penalty ¹.

¹Calvetti D, Pascarella A, Pitolli F, Somersalo E, Vantaggi B (2015) A hierarchical Krylov–Bayes iterative inverse solver for MEG with physiological preconditioning. Inverse Problems 31:125005

A convergence result

Let the function for updating the variance $f: \mathbb{R}^n \to \mathbb{R}^n$ have components

$$f_j(x_j) = heta_j^*\left(rac{\eta}{2} + \sqrt{rac{\eta^2}{4} + rac{x_j^2}{2 heta_j^*}}
ight)$$

Theorem

For a Gamma hyperprior, the exact IAS algorithm converges to the unique minimizer $(\hat{x}, \hat{\theta})$ of the Gibbs energy functional. Moreover, the minimizer $(\hat{x}, \hat{\theta})$ satisfies the fixed point condition

$$\widehat{x} = \operatorname{argmin} \left\{ \mathscr{E}(x \mid F(x)) \right\}, \quad \widehat{\theta} = F(\widehat{x}),$$

where F is the map with jth component f_i ².

²Calvetti D, Pascarella A, Pitolli F, Somersalo E, Vantaggi B (2015) A hierarchical Krylov–Bayes iterative inverse solver for MEG with physiological preconditioning. Inverse Problems 31:125005

Scale parameter and sparsity

Under the assumptions of our hierarchical Bayesian model we have shown that

• The exact IAS iteration converges to the global minimizer of the functional

$$\mathscr{L}_{\eta}(x) = \mathscr{E}(x, f(x))$$

and, for small $\eta > 0$

$$\mathscr{L}_{\eta}(x) = \mathscr{L}_{0}(x) + \underbrace{\eta g(x, \eta)}_{\to 0 \text{ as } \eta \to 0},$$

where

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$$\mathscr{L}_0(x) = rac{1}{2} \|b - \mathsf{A}x\|^2 + \sqrt{2} \sum_{j=1}^n rac{|x_j|}{\sqrt{ heta_j^*}}.$$

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Gamma hyperprior parameters

From the result above it follows that

- For small η , the IAS minimization problem is a small perturbation of the weighted ℓ_1 penalized least squares functional
- **2** The parameter η controls the sparsity of the solution.
- O The scale parameters θ^{*}_j play the role of sensitivity weights in inverse problems
- Oata components may have different sensitivity to different components x_i.

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ℓ_2 Stable Signal Recovery

Two remarks $\underbrace{x_{\eta} = \operatorname{argmin} \left\{ \mathscr{L}_{\eta}(x) \right\}}_{=IAS \text{ solution}} \qquad \underbrace{x_{1} = \operatorname{argmin} \left\{ \mathscr{L}_{0}(x) \right\}}_{=\ell_{1} \text{ penalized solution}}.$

() The difference $x_{\eta} - x_1$ is a vector whose size depends continuously on η .

② If A has the Restricted Isometry Property (RIP) and the data comes from a sparse vector³, then x_{η} is close to the underlying sparse solution.

³Candes E, Romberg JK and Tao T(2006): Stable Signal Recovery from Incomplete and Inaccurate Measurements, Comm Pure Appl Math LIX: 1207–1223

Sparse signal and exchangeability

Assume the underlying signal x is sparse $supp(x) = I \subset \{1, 2, ..., n\}$ and b_0 is the noiseless measurement. Define

$$SNR_{|_{I}} = \frac{E\left\{\|b_{0}\|^{2} \mid supp(x) = I\right\}}{E\left\{\|e\|^{2}\right\}}.$$

Lemma

With our assumptions about X and E

$$\mathrm{SNR}_{|_{I}} = rac{\sum_{j \in I} eta heta_{j}^{*} \|\mathsf{A} \mathbf{e}_{j}\|^{2}}{\mathrm{tr}(\Sigma)}.$$

Proof.

$$\textit{E}\left\{ \|\textit{b}_{0}\|^{2}\right\} = \mathrm{Tr}\textit{E}\left\{\textit{b}_{0}\textit{b}_{0}^{\mathrm{T}}\right\} = \mathrm{Tr}\textit{E}\left\{\mathsf{Axx}^{\mathrm{T}}\mathsf{A}^{\mathrm{T}}\right\} = \mathrm{Tr}\left(\mathsf{A}\textit{E}\left\{\mathsf{xx}^{\mathrm{T}}\right\}\mathsf{A}^{\mathrm{T}}\right),$$

and from the Gamma hyperprior

$$E\left\{xx^{\mathrm{T}}\right\} = E_{\theta}\left\{E\left\{xx^{\mathrm{T}} \mid \theta\right\}\right\} = E\left(\mathrm{diag}(\theta)\right) = \mathrm{diag}(\beta\theta^{*}).$$

Choice of scale parameter

How should θ^* be chosen?

Theorem

Given an estimate $\overline{\mathrm{SNR}}$ of SNR, if

$$P(||x||_0 = k) = p_k, \quad p_0 = 0, \quad \sum_{k=1}^n p_k = 1$$

and if

$$\mathrm{SNR}\mid_I = \mathrm{SNR}\mid_{I'}, \quad \forall \ I, I': \mathrm{card}(I) = \mathrm{card}(I'),$$

then

$$heta_j^* = rac{\mathsf{C}}{\|\mathsf{A} e_j\|^2}, \quad \mathsf{C} = \overline{\mathrm{SNR}} \mathrm{Tr} ig(\Sigma ig) \sum_{j=1}^n rac{p_k}{k}$$

In the literature $||Ae_j||^2$ is the *sensitivity* of the data to *j*th component of *x*.

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Sparsity and quadratic convergence

Theorem

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- $f_j(x_j)$ is monotonically increasing
- $f_j(0) \geq \tilde{\theta} > 0$
- x_* is sparse with $\operatorname{supp}(x_*) = T$, |T| = S

then exact IAS converges quadratically in θ_{T^C} , where T^C is the complement of T.

Theorem

Under the conditions of the previous theorem, if

- *x*_{*} is nearly sparse (compressible)
- $\|x_T c\|_{\infty} \leq \zeta$

then

$$J_{T^c} \leq \sqrt{S} \frac{\zeta}{\tilde{ heta}}.$$

Sparsity and quadratic convergence

The theorems implies that

- If ζ is small enough, the convergence is effectively quadratic in θ_{T^C}
- $\bullet~$ The ℓ_1 solution from noisy data approximates well the underlying sparse signal 4,
- As $\beta \to (3/2)^+$, the exact IAS solution approaches the ℓ_1 penalized solution hence is close to the underlying sparse signal.

⁴Candes E, Romberg JK and Tao T(2006): Stable Signal Recovery from Incomplete and Inaccurate Measurements, Comm Pure Appl Math LIX: 1207–1223

6. Inexact IAS

Inexact IAS and quasi-MAP estimate

● For large scale problems and few observations (A ∈ ℝ^{m×n}, m < n), the least squares step of IAS</p>

$$\begin{bmatrix} \mathsf{A} \\ \mathsf{D}_{\theta}^{-1/2} \end{bmatrix} \mathbf{x} = \begin{bmatrix} \mathbf{b} \\ \mathbf{0} \end{bmatrix}$$

can be solved approximately by priorconditioned CGLS⁵ ⁶ If $\Sigma^{-1} = S^{T}S$, apply the CGLS method to

$$\mathsf{SAD}_{\theta}^{1/2}w = \mathsf{S}b$$

stopping on Morozov discrepancy principle at *j*th step, where

$$d_j > \sqrt{m} > d_{j+1}, \quad d_j = \|\mathsf{S}b - \mathsf{SAD}_\theta^{1/2}w_j\|.$$

S Retrieve original variable $x_k = D_{\theta}^{1/2} w_j$

 5 Calvetti et al. Priorconditioned CGLS-Based Quasi-MAP Estimate, Statistical Stopping Rule, and Ranking of Priors. SIAM J. Sci. Comput. 39-5 (2017)

Pros of inexact IAS

- Suitable for large problems
- Omputationally efficient
- Sollows classical scheme of inner/outer iterations
- Can be interpreted as flexible right preconditioning
- Sumber of CGLS steps decreases with outer iterations
- When the underlying signal is sparse, the algorithm automatically reduced the effective dimensionality of the problem to solve

Image: A matrix

Open questions for inexact IAS

At the present, the inexact IAS

- Does not have a proof of convergence, only numerical evidence
- The solution produced is an approximation of the MAP estimate, hence we call it quasi MAP estimate (qMAP)
- A quadratic rate of convergence of signal and prior variances has been observed numerically: a rigorous proof is in progress.

Computed examples

- $\bullet~A \in 1000 \times 1000$ is a Gaussian blurring kernel over 15 pixels
- Data come from a piecewise constant signal
- Additive scaled white noise 0.01% of max of noise-free data
- Conditionally Gaussian first order smoothness prior
- Gamma hyperprior

Image: A matrix and a matrix



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Signal Estimate



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An application to MEG

In this example we see the effect of the focality parameter

- Data: 153 measurements at magnetometers
- \bullet A is the leadfield matrix 153×75000
- Sparsity prior

Image: A mathematical states and a mathem

8. Computed examples



Hierarchical Bayesian Sparsity: ℓ_2 Magic