Fast Point Cloud Distances and Multi-Sample Testing

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Outline

1 Classical Statistics Two Sample Problem

- 2 Reframing as an Optimization Problem
- Interplay with Analysis and Spectral Theory
- 4 Bounding Quadrature Error
- 5 Data Science and Medical Applications

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• **Question:** Suppose $X \sim p$ and $Y \sim q$ on \mathbb{R}^d

$$egin{array}{ll} H_0:&p=q\ H_1:&p
eq q \end{array}$$

- Statistical in nature
 - · Goal is to convince you this touches on
 - Spectral theory
 - Optimization
 - Data science in medicine

Two Sample Test Applications

• **Question:** Suppose $X \sim p$ and $Y \sim q$ on \mathbb{R}^d

$$\begin{array}{ll} H_0: & p = q \\ H_1: & p \neq q \end{array}$$



Two Sample Tests in 1D

• **Question:** Suppose $X \sim p$ and $Y \sim q$ on \mathbb{R}^d

$$\begin{array}{ll} H_0: & p=q\\ H_1: & p\neq q \end{array}$$

Answer:

• Easy in 1D: Kolmogorov-Smirnov



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1	0.950	0.975	0.995		
2	0.776	0.842	0.929		
3	0.642	0.708	0.818		
4	0.564	0.624	0.735	1.000	1.000
5	0.510	0.565	0.669	1.000	1,000
6	0.418	0.521	0.618	0.833	1,000
7	0.458	0.486	0.577	0.857	0.817
- 8	0.411	0.457	0.543	0.750	0.875
9	0.388	0.432	0.514	0.667	0.778
10	0.368	0.410	0,490	0.700	0.800
- 11	0.352	0.391	0.468	0.636	0.527
12	0.338	0.375	0.450	0.583	0.667
13	0.315	0.361	0.433	0.538	0.692
14	0.314	0.349	0.418	0.571	0.645
14	0.104	0.118	0.404		0.600
16	0.295	0.328	0.392	0.500	0.625
17	0.286	0.318	0.381	0.471	0.588
18	0.778	0.509	0.371	0.500	0.556
19	0.272	0.321	0.363	0.474	0.516
70	0.764	0.294	0.356	0.450	0.510
	0.74	0.27	0.37	0.40	0.48
30	0.77	0.74	0.29	0.37	0.43
35	0.21	0.23	0.27	0.34	0.39
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• Hard in nD: Topic of this work

Two Sample Tests in nD

- Why is it hard in higher dimensions:
 - Marginals of distribution are insufficient



- Difficult to define relevant bins
- Curse of dimensionality: most bins will have very few points
 - Minimax rate exhibits curse (Arias-Castro et al 2017)



Two Sample Tests in nD

• More general question: How do we define a distance between *X* and *Y*?

$$d(X, Y) < \epsilon \implies p = q$$

 $d(X, Y) > \epsilon \implies p \neq q$

Additional Questions

- Not just interested in whether they deviate, but how and where?
- Can we extend to *k*-samples and use distance matrix between pairwise samples?
- How without assumption of underlying manifold structure?

Classical Statistics Two Sample Problem

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Bins and Locality

- No point-to-point correspondence, so use bins/histograms
 - Bins too small, lead to high variance
 - Bins too large, lead to poor precision
- More generally by maximum mean discrepancy with some set of functions

$$MMD(p,q;\mathcal{F}) = \sup_{f\in\mathcal{F}} \left(\int f(x)dp(x) - \int f(x)dq(x) \right)$$

$$\widehat{MMD}(X,Y;\mathcal{F}) = \sup_{f\in\mathcal{F}} \left(\frac{1}{|X|} \sum_{x\in X} f(x) - \frac{1}{|Y|} \sum_{y\in Y} f(y) \right)$$



Avoid Optimization with Kernels

- Problem: Which function classes are tractable?
- **Possible Solution:** Bins defined by a *kernel* $k : \mathbb{R}^d \times \mathbb{R}^d \to [0, 1]$

$$k(x, y) = e^{-||x-y||^2/\sigma^2}$$
 (Example)

• Take \mathcal{F} as unit ball in Reproducing Kernel Hilbert Space $\mathcal{H}(k)$

$$f \in \mathcal{H} \text{ if } \mathcal{F}_t[f] = \langle f, k(t, \cdot) \rangle = f(t)$$





Example Balls of Affinity > 0.1

Maximum Mean Discrepency

• Want to define density at any $z \in \mathbb{R}^d$

$$\mu_{p}(z) = \mathbb{E}_{x \sim p}[k(z, x)]$$

Then avoid optimization

$$MMD^{2}(p, q; \mathcal{F}) = \left[\sup_{f \in \mathcal{F}} \left(\mathbb{E}_{x}f(x) - \mathbb{E}_{y}f(y)\right)\right]^{2}$$
$$= \left[\sup_{f \in \mathcal{F}} \langle \mu_{p} - \mu_{q}, f \rangle\right]^{2}$$
$$= \|\mu_{p} - \mu_{q}\|_{\mathcal{H}}^{2}$$



Figure 1.1: Embedding of marginal distributions: Each distribution is mapped into an RKHS via an expectation operation.

Skölkoph

Discrete MMD

Still hard to compute, but exists equivalent discrete version

$$MMD(X, Y) = \frac{1}{n(n-1)} \sum_{x, x' \in X} k(x, x') + \frac{1}{m(m-1)} \sum_{y, y' \in Y} k(y, y') - \frac{2}{mn} \sum_{x \in X, y \in Y} k(x, y)$$

- Avg. Affinity within *X*, Avg. Affinity within *Y*
- Avg. Affinity between X and Y







Guarantees (Gretton, et al. 2011)

- $\mathit{MMD}(p,q)$ is one-to-one mapping for $\|p-q\|_\infty$
- If p = q, $MMD(X, Y) \rightarrow 0$ like $\frac{C}{\sqrt{n+m}}$
 - C depends on bandwidth of kernel
- If $p \neq q$, *MMD*(*X*, *Y*) minimum distance detectable is $\|\mu_q \mu_p\| = \frac{c}{\sqrt{n+m}}$

Problems

- Kernel is isotropic
 - Treats data on single scale
 - Convergence depends on dimension of ambient space (Wasserman et al 2014)
- $p \neq q$ results don't speak to power of test
- $O(n^2)$ storage of K
 - Completely intractable for *k*-sample problem and network geometry

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- $O(n^2)$ storage of K
 - Completely intractable for *k*-sample problem and network geometry
- Introduction of local geometry creates data-adaptive test
 - Penalizes moving in normal direction
 - Discounts regions of high-volatility
 - Allows assumptions of local low dimensionality and off manifold deviation





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Spectral Convergence

• Easier to work with mean centered kernel \tilde{k} and spectrum

$$ilde{k}(x,y) = \sum_{i} \sigma_i \phi_i(x) \phi_i(y), \quad ilde{k} \in L^2(\mathcal{X} \times \mathcal{X}, p \times p), \quad \sigma_i \to 0$$

- Eigenfunctions have a long history of determining shape
 - Principle Component Analysis
 - Spectral graph theory
- As note, $(D(k) k) \phi = \lambda \phi \rightarrow -\Delta f = \lambda f$
 - Implies eigenvalues of \tilde{k} not sufficient to differentiate between datasets



Convergence of Eigenvectors

• If p = q, goal to show

$$rac{1}{|X|}\sum_{x\in X}\phi_i(x)
ightarrow 0, \qquad rac{1}{|Y|}\sum_{y\in Y}\phi_i(y)
ightarrow 0$$

• If $p \neq q$, show convergence to non-zero constants

- Must treat ϕ_i simultaneously b/c samples not independent
 - Multi-dim central limit theorem and spectral decay



Eigenvectors are "Redundant"

- Approximate rank of *K* determined by number of balls needed to cover data (Tygert, Rokhlin 2008; Kühn 2011)
 - For acceptable bandwidth, $|\{i|\sigma_i(K) > \epsilon\}| \ll n$
- Using square matrix is over-redundant
 - Can choose small set of "reference points" with balls that easily cover data

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Anisotropic Kernels

- Assume ambient space has local geometry (r, Σ_r)
 - Sample r from $X \cup Y$
 - Random subsample
 - QR with pivoting
 - Quadrature methods
 - Σ_r from covariance of nearest neighbors
- Choose set of "representative points" R to compare distributions
 - Test only as good as reference points

$$A(r,x) = e^{(x-r)^{\mathsf{T}} \sum_{r=1}^{n-1} (x-r)}$$

• Still have mean embedding $\mu_X(r) = \mathbb{E}_{x \in X}[A(r, x)]$



Adaptive Kernel MMD

• $AMMD(X, Y; \{r, \Sigma_r\}) = \|\mu_X(r) - \mu_Y(r)\|_2$

- Distance of distribution proj. to lighthouses
- Lighthouses contain directional and dimension information
- Calculation is O(N|R|d)
 - Never calculating eigendecomposition



Assumption: R sufficiently span column space of K

- Parallel to kernel "having non-vanishing Fourier transforms on any interval" for KDE
- $q = p + \tau g$

Informal Theorems (C., Cheng, Coifman 2017)

All shifts and variance depend on spectrum and how quickly $\int \psi_k(y)g(y)dy$ detect deviation

• If
$$\tau = O(n^{-1/2})$$
, then nT_n is χ^2

- ② If $\tau = O(n^{-1/2+\delta})$, then $n^{1-\delta}T_n$ is normal with shift "larger than variance" as *n* → ∞
- If $\tau = 1$, $\sqrt{n}T_n$ is normal with shift "larger than variance"

- Need to know if deviation is enough in comparison to H₀
 - Either true H₀ or permutation null
- Important to know not just threshold but for fixed deviation
 - Power is prob to detect deviation when it exists

Power (C., Cheng, Coifman 2017)

Notations as above, under Assumption, for specific g = q - p fixed, if $\tau_n = O(n^{-1/2+\delta})$ where $0 < \delta \le \frac{1}{2}$ ($\delta = \frac{1}{2}$ means that $\tau = 1$), the test power $\pi_n(p + \tau_n g) \to 1$ as $n \to \infty$.

Also extends for MMD results of Gretton et al



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Bounded Diffusion Subsampling

- Choosing (r, Σ_r) currently done by random sampling of X ∪ Y
 - Effectively Monte Carlo integration
- Reframe as problem of efficiently learning eigenfunction means in CLT

Bounded Diffusion Subsampling

- Choosing (r, Σ_r) currently done by random sampling of X ∪ Y
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Near Perfect Spherical Designs; (Steinerberger, Linderman, 2018)

For $a_r \ge 0$ and $\sum a_r = 1$, then for any *f* expressible in terms of low-freq eigenfunctions of graph G = (X, K),

$$\left|\frac{1}{|R|}\sum_{r\in R}a_rf(r)-\mathbb{E}_{x\in X}[f(x)]\right|\leq \frac{\|f\|_{X_{\lambda}}}{\lambda^t}\left(\left\|K^t\sum_{r\in R}a_r\delta_r\right\|_2^2-\frac{1}{|X|}\right)^{1/2}$$

Geometric interpretation: *R* distributed to diffuse sufficiently quickly for small *t*.

Reference points subsample $\int_{z \in X \cup Y} f(z) dz = \int_{z \in X \cup Y} |\mu_X(z) - \mu_Y(z)|^2 dz$

Non-Asymptotic MMD Reference Error; (C. 2018)

Assuming $\mu_X - \mu_Y$ has energy τ and projects mostly projects "mostly" $(1 - \epsilon)$ onto eigenfunctions $\lambda > \nu$, for fixed reference set *R*,

$$\left|AMMD(X, Y; a_r) - MMD(X, Y; K)\right|^2 < \frac{\tau^2}{\nu^2} \left(\left\|K \sum_{r \in R} a_r \delta_r\right\|_2 - \frac{1}{N} \right) + \epsilon^2 \tau^2 \left(\sum_{\lambda < \nu} \lambda^2\right)$$

Effectively dominated by finite error $\frac{c}{\sqrt{N}}$ Currently working on fast preprocessing optimization scheme to minimize over choice of *R* and *a*_r



- For k-sample problem, only need comparison to fixed reference set
- Requires one loop through data for AMMD, and one loop over $\mu_{\rho}(r)$

$$k-\text{sample test from } O\left(\binom{K}{2}N^2 \cdot D\right) \quad \text{to } O\left(\binom{K}{2}|R| + K \cdot N|R|D\right)$$

IsotropicMMD RefSetAMMD

Using MMD

- In practice, can use permutation test
 - Set $Z = X \cup Y$ and define permutation $p = \begin{bmatrix} p_1 & p_2 \end{bmatrix}$ of $\{1, ..., n + m\}$
 - **2** Compute $MMD(Z_{p_1}, Z_{p_2})$
 - 3 Create histogram over N iterations of *p* for null hypothesis
 - **(4)** If $\#\{p: MMD(X, Y) > MMD(Z_{p_1}, Z_{p_2})\}/N > 0.95$, reject H_0



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Comparison to Isotropic Gaussian

• At the moment, mostly empirical comparison of spectrum and deviation



Adaptive Kernels Example



Figure: Estimated $\xi_p(r)$ and $\xi_q(r)$ for two distributions *p* and *q* with 1000 samples each (left) with gaussian kernel (middle) and anisotropic kernel (right) respectively.

Real Example

Flow cytometry: each patient is represented by 9D point cloud of cells



• Used to tell if people have blood disease

Medical test is to look at every 2D slice



Healthy

AML

Unsupervised Organization:

- Measure distance between every two people X_i and X_j
 - Form network of people
 - Use eig. fctns. of distance matrix for unsupervised clustering.

Distance Between Flow Cytometry Measurements



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Interpretability and Classification

Supervised Information:

Kernel MMD has dual form

$$MMD(p, q, \mathcal{F}) = \sup_{f \in \mathcal{F}} |\mathbb{E}_{x \in p} f(x) - \mathbb{E}_{y \in q} f(y)|$$

- Arg max f is called witness function
 - Interprets regions of "importance" to differentiate p and q
- Computable as

$$f(x) = \int \left(\int a(r, x)a(r, y)p_R(r)dr\right)(p(y) - q(y))dy$$



Flow Cytometry from MDS

- MDS is a more general class of blood cancers
- Much more difficult to detect than AML





Permutation Test Anisotropic Kernel

General Take-Aways

- Adaptive kernels are necessary tool for data science
- Kernels allow extreme flexibility for defining point-to-point correspondence
 - Multi-scale affinity kernels
 - Function model driven kernels
 - Features generated from any transform

$$k(x,y) = e^{-\|\Phi(x) - \Phi(y)\|^2/\sigma^2}$$

- Extensions:
 - Multi-bandwidth generalization
 - Local significance of deviation
 - Generalization to non-iid samples (time series)
 - GANs

Questions?