

# Bayesian static parameter estimation using Multilevel Monte Carlo

**Kody Law**

joint with A. Jasra (NUS), K. Kamatani (Osaka), & Y. Zhou

MANCHESTER  
1824

The University of Manchester

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**Aim:** Approximate expectations with respect to a target probability distribution  $\eta_\infty$ , which needs to be approximated by some  $\eta_L$ , and the un-normalized target can only be evaluated up to a non-negative unbiased estimator.

**Solution:** Apply an approximate coupling strategy so that the multilevel Monte Carlo (MLMC) method can be used within a pseudo-marginal algorithm [B02, AR08, ADH10].

- MLMC methods *reduce cost to error* =  $\mathcal{O}(\varepsilon)$  [G08];
- Recently this methodology has been applied to *inverse UQ*, mostly in cases where  $\eta_L$  **can be evaluated** up to a normalizing constant [HSS13, DKST15, HTL16, BJLTZ17].
- Here it is assumed that we **cannot**,
- however we can construct a **non-negative and unbiased estimator** of the un-normalized target  $\gamma_L$  [JKLZ18].



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# Example: expectation for SDE [G08]

Estimation of expectation of solution of intractable stochastic differential equation (SDE).

$$dX = f(X)dt + \sigma(X)dW, \quad X_0 = x_0.$$

**Aim:** estimate  $\mathbb{E}(g(X_T))$ .

We need to

- (1) Approximate, e.g. by Euler-Maruyama method with resolution  $h$ :

$$X_{n+1} = X_n + hf(X_n) + \sqrt{h}\sigma(X_n)\xi_n, \quad \xi_n \sim N(0, 1).$$

- (2) Sample  $\{X_{N_T}^{(i)}\}_{i=1}^N$ ,  $N_T = T/h$ .

# Assumptions

**(A1)**  $\forall y \in \mathcal{T}, \exists C > 0$  such that  $\forall x \in \mathcal{S}, \theta \in \Theta,$

$$C \leq g_{\theta}(x, y) \leq C^{-1}.$$

And  $\forall y \in \mathcal{T}, g_{\theta}(x, y)$  is globally Lipschitz on  $\mathcal{S} \times \Theta$ .

**(A2)**  $\forall 0 \leq k \leq n, q \in \{1, 2\}, \exists \beta > 0$  such that  $\forall$   
 $\varphi \in \mathcal{B}_b(\Theta \times \mathcal{S}^{k+1}) \cap \text{Lip}(\Theta \times \mathcal{S}^{k+1}) \exists C > 0$

$$\left( \int_{\Theta \times \mathcal{S}^{2k+2}} |\varphi(\theta) - \varphi(\theta)|^q \Pi(d\theta) \nu_{\theta}(dz_0) \prod_{p=1}^k Q_{\theta, h, h'}(z_{p-1}, dz_p) \right)^{3-q} \leq C(h')^{\beta}.$$

**(A3)** Suppose that  $\forall n > 0, \exists \xi \in (0, 1)$  and  $\nu \in \mathcal{P}(\mathcal{W})$  such that  
for each  $w \in \mathcal{W}, \varphi \in \mathcal{B}_b(\mathcal{W}) \cap \text{Lip}(\mathcal{W}), h, h'$ :

$$\int_{\mathcal{W}} \varphi(w') K(w, dw') \geq \xi \int_{\mathcal{W}} \varphi(v) \nu(dv).$$

$K$  is  $\eta$ -reversible, that is,

$$\int_{\mathcal{W} \in B} \eta(dw) K(w, A) = \int_{\mathcal{W} \in A} \eta(dw) K(w, B) \text{ for any } A, B \in \mathcal{W}.$$

## Theorem (JKLZ18)

*Assume (A1-3). Then  $\forall n > 0, \exists \beta > 0$  such that  $\forall \varphi \in \mathcal{B}_b(\Theta \times \mathbb{S}^{n+1}) \cap \text{Lip}(\Theta \times \mathbb{S}^{n+1}) \exists C > 0$  such that*

$$\mathbb{E}[\bar{E}_I^{N_I}(\varphi)^2] \leq \frac{Ch_I^\beta}{N_I}.$$

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(A4)  $\exists \gamma, \alpha, C > 0$  such that the cost to simulate  $E_l^{N_l}$  is controlled by  $C(E_l^{N_l}) \leq CN_l h_l^{-\gamma}$ , and the bias is controlled by

$$|\mathbb{E}_{\Pi_{h_L}}(\varphi(\theta, X_{0:n})) - \mathbb{E}_{\Pi_0}(\varphi(\theta, X_{0:n}))| \leq Ch_L^\alpha.$$

## Corollary

Assume (A1-4).  $\forall n > 0$  and  $\varphi \in \mathcal{B}_b(\Theta \times \mathbb{S}^{n+1}) \cap \text{Lip}(\Theta \times \mathbb{S}^{n+1}) \exists C > 0$  such that  $\forall \epsilon > 0$  one can choose  $(L, \{N_l\}_{l=1}^L)$  so

$$\mathbb{E} \left[ \left| \sum_{l=0}^L E_l^{N_l}(\varphi) - \mathbb{E}_{\Pi_0}(\varphi(\theta, X_{0:n})) \right|^2 \right] \leq C\epsilon^2,$$

with a total cost (per time step)

$$\text{COST} \leq C \begin{cases} \epsilon^{-2}, & \text{if } \beta > \gamma, \\ \epsilon^{-2} |\log(\epsilon)|^2, & \text{if } \beta = \gamma, \\ \epsilon^{-(2 + \frac{\gamma - \beta}{\alpha})}, & \text{if } \beta < \gamma. \end{cases}$$



$$dX_t = \frac{1}{2} \nabla \log \pi(X_t) dt + \sigma dW_t, \quad X_0 = x_0$$

$$Y_k | X_k \sim \mathcal{N}(0, \tau^2 \exp X_k),$$

$$\theta \sim \mathcal{G}(1, 1),$$

$$\sigma \sim \mathcal{G}(1, 0.5).$$

- $\pi(x)$  denote the probability density function of a Student's  $t$ -distribution with  $\theta$  degrees of freedom.
- $x_0 = 0$ .
- 1,000 observations simulated with  $\theta = 10$ ,  $\sigma = 1$ , and  $\tau^2 = 1$ .

Parameter  $\sigma$   $\theta$

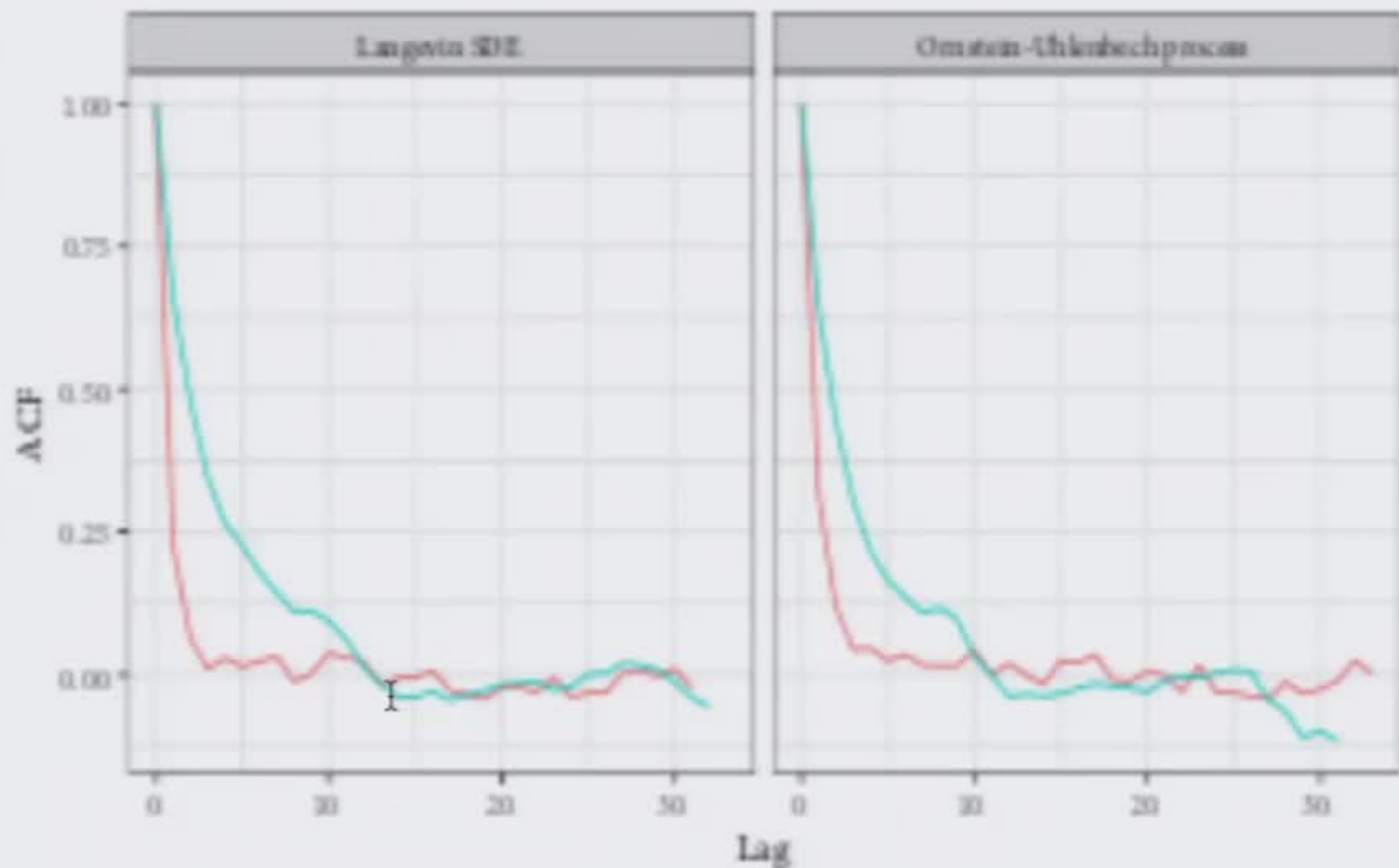


Figure: Autocorrelation of a typical PMCMC chain.

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Algorithm ● ML-PMCMC ● PMCMC

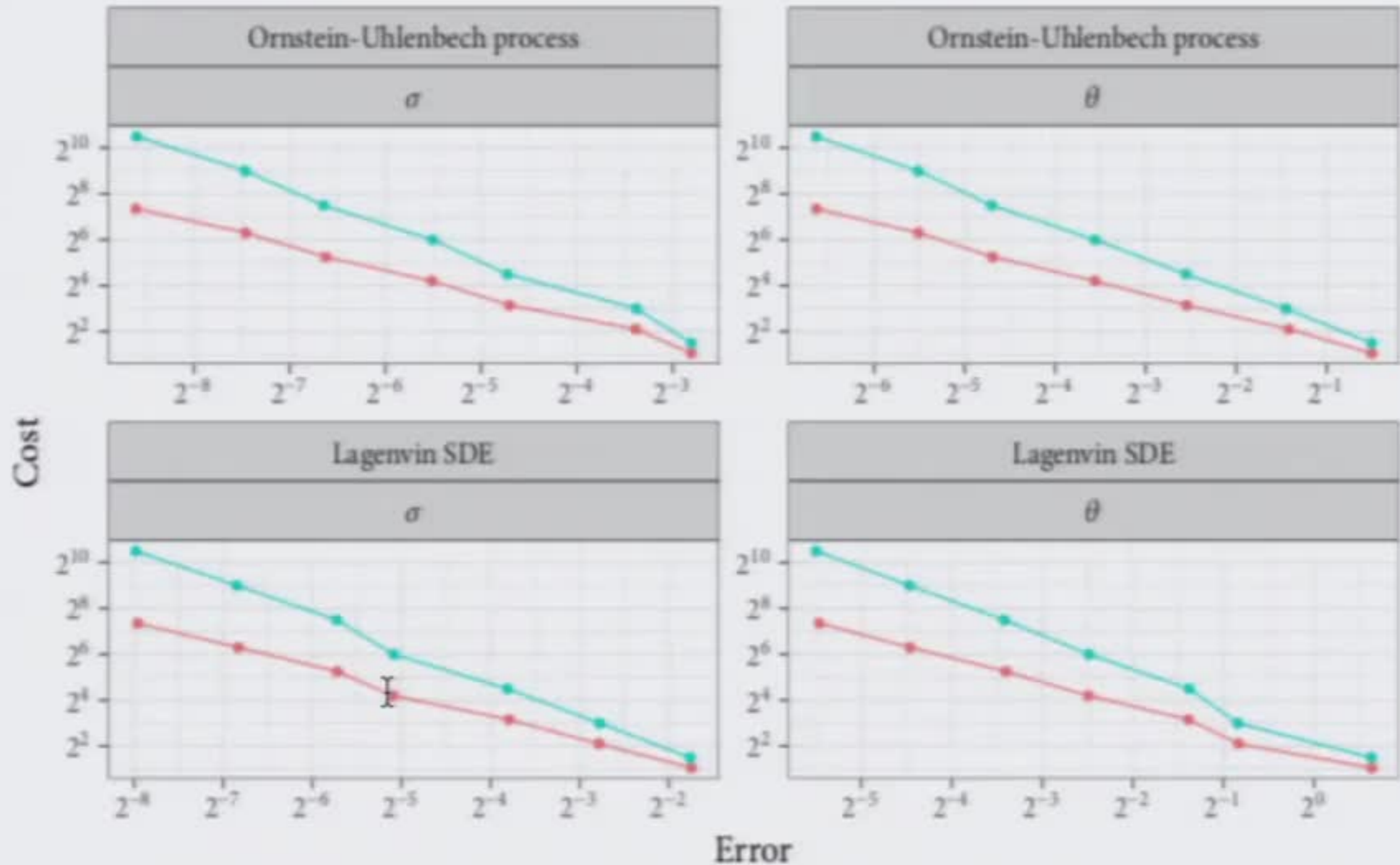


Figure: Cost vs. MSE for the 2 parameters for each of the 2 SDEs.

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Model	Parameter	ML-PMCMC	PMCMC
OU	$\theta$	-1.022	-1.463
	$\sigma$	-1.065	-1.522
Langevin	$\theta$	-1.060	-1.508
	$\sigma$	-1.023	-1.481

**Table:** Estimated rates of convergence of MSE with respect to cost for various parameters, fitted to the curves.

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- New approximate coupling strategy can be used to apply MLMC to PMCMC for static parameter estimation [JKLZ18.i].
- Same strategy can be employed for multi-index MCMC [JKLZ18.ii].
- In progress: SMC<sup>2</sup> [JLX18+]

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